

## STA447/2006 HW#1 Solutions – OUTLINE ONLY

NOTE: What follows is just an outline of the homework solutions, taken mostly from the textbook's solutions manual. These solutions may be incorrect or incomplete; more complete explanations are required to earn full points on the homework.

**9.4.** (a) to go from 1 to 4 in three steps we must go 1,2,3,4 so  $p^3(1,4) = (.4)^3 = .064$ . (b) to go from 1 to 0 in three steps we may go 1,2,1,0 or 1,0,0,0 so  $p^3(1,0) = (.4)(.6)^2 + .6 = .744$

**9.8.** (a)  $\{1, 5\}$  and  $\{2, 5\}$  are finite irreducible closed sets so all of these states are recurrent.  $3 \rightarrow 1$  but  $1 \not\rightarrow 3$  so 3 is transient.  
 (b)(a)  $\{1, 4\}$  and  $\{2, 5\}$  are finite irreducible closed sets so all of these states are recurrent.  $3 \rightarrow 2$  but  $2 \not\rightarrow 3$  so 3 is transient.  $6 \rightarrow 1$  but  $1 \not\rightarrow 6$  so 6 is transient.

**Question 3:** (a) States 1,6,7 are recurrent; states 2,3,4,5 are transient. (b) [Can use e.g. the expansion that  $f_{ij} = p_{ij} + \sum_{k \neq j} p_{ik} f_{kj}$ .] Answers are  $f_{11} = f_{21} = f_{31} = 1$ ,  $f_{61} = f_{71} = 0$ ,  $f_{51} = 1/3$ , and  $f_{41} = (1/2) + (1/2)(1/3) = 2/3$ .

**9.1.** No. We first argue this intuitively: when  $X_n = 1$  the last two results may be 0, 1 or 1, 0. In the first case we may jump only to 2 or 1, while in the second case we may only jump to 0 or 1. Thus it is not enough to know just current state. To get a formal contradiction we note that  $X_1 = 2$ ,  $X_2 = 1$  implies that  $Y_0 = Y_1 = 1$ ,  $Y_2 = 0$  so

$$P(X_3 = 2 | X_2 = 1, X_1 = 2) = 0 < P(X_3 = 2 | X_2 = 1)$$

**9.36.** (a) To increase the number of black balls by 1 we have to pick a white from the first and a black from the second. Using similar reasoning on the other cases we have

$$p(i, i+1) = \frac{(m-i) \cdot (b-i)}{m \cdot m} \quad p(i, i-1) = \frac{i \cdot (m-b+i)}{m \cdot m}$$

$$p(i, i) = \frac{i(b-i) + (m-i)(m-b+i)}{m \cdot m}$$

(b) Detailed balance is equivalent to  $\pi(i+1)/\pi(i) = p(i, i+1)/p(i+1, i)$ . To check this condition we compute

$$\begin{aligned} \frac{\pi(i+1)}{\pi(i)} &= \frac{b!}{(i+1)!(b-i-1)!} \frac{(2m-b)!}{(m-i-1)!(m-b+i+1)!} \\ &\quad \cdot \frac{i!(b-i)!(m-i)!(m-b+i)!}{b! (2m-b)!} \\ &= \frac{i!(b-i)!(m-i)!(m-b+i)!}{(i+1)!(b-i-1)!(m-i-1)!(m-b+i+1)!} \\ &= \frac{(b-i)(m-i)}{(i+1)(m-b+i+1)} = \frac{p(i, i+1)}{p(i+1, i)} \end{aligned}$$

(c) If we pick  $m$  balls at random from the  $2m$  to put in one urn then  $\pi(i)$  gives the probability we will get exactly  $i$  black balls. If we put the balls in randomly and then switch two we still have a random arrangement so we have constructed a stationary distribution.

**9.32.** (a)  $E_x X_1 = x + (N-x)/N - x/N = 1 + (1-2x)/N$ . Taking  $x$  to be random with distribution equal to that of  $X_n$  the result follows. (b) The formula holds when  $n=0$ . If it is true for  $n$  then

$$\begin{aligned} \mu_{n+1} &= 1 + (1-2/N) \left( \frac{N}{2} + (1-2/N)^n (x - N/2) \right) \\ &= \frac{N}{2} + \left( 1 - \frac{2}{N} \right)^{n+1} (x - N/2) \end{aligned}$$

**9.31.** (a) When  $u, v > 0$  the chain is irreducible so the conclusion follows from (4.7) and (4.5). (b) By the formula for the mean of the binomial  $E_x X_1 = N\rho_x = vN + (1 - u - v)x$ . Iterating we have

$$\begin{aligned} E_x X_2 &= vN + (1 - u - v)vN + (1 - u - v)^2 x \\ E_x X_3 &= vN + (1 - u - v)vN + (1 - u - v)^2 vN + (1 - u - v)^3 x \end{aligned}$$

so  $\lim_{n \rightarrow \infty} E_x X_n = vN \sum_{m=0}^{\infty} (1 - u - v)^m = Nv/(u + v)$ . A simple way to see this is to note that in the absence of mutation an individual in generation  $n$  is the same as a randomly chosen one from generation  $n - 1$ , so its type is dictated by the first mutation encountered as we work backwards and that will be to a 1 with probability  $v/(u + v)$ .

**9.35.** Let  $j$  be a possible arrangement of the deck. There are exactly 52 arrangements  $k$  that can arise from  $j$  in one step. In the other direction there are 52 arrangements  $i$  that can lead to  $j$  in one step. Since each transition has probability  $1/52$ , it follows that  $\sum_i p(i, j) = 52 \cdot (1/52) = 1$ . The matrix is doubly stochastic, so by a result in class the stationary distribution is uniform.

(You also have to prove the chain is irreducible and aperiodic, though that is not too hard.)